

An empirical model of the decision to switch between electricity contracts

Gauthier Lanot Mattias Vesterberg

Department of Economics
Umeå School of Business and Economics,
Umeå University, Sweden

Oslo, 23 April 2018

Active consumers

Recently, much policy focus on consumers as active participants on the electricity market

- ▶ Integration of renewable energy
- ▶ Smart grids, demand flexibility
- ▶ Price responsiveness is key

Many margins of choice

On a de-regulated electricity market, households face many (complex) margins of choice:

- ▶ Appliances/technology/efficiency
- ▶ Utilization of appliances
- ▶ Choice of retailer
- ▶ Choice of transmission capacity
- ▶ Choice of electricity contract (focus of this paper!)

Aggregate over all these decisions determine electricity demand, with obvious policy implications: e.g., demand flexibility.

This paper:

In Sweden, two common contract types:

- ▶ price per kWh vary by month ("variable-price contract")
- ▶ price per kWh fixed for a year or longer ("fixed-price contracts")

We focus on households on variable-price contracts deciding either to remain on this contract, or switch to a fixed-price contract in response to prices.

- ▶ A household choosing a variable-price contract is assumed to believe this to be the most preferable contract.
- ▶ However, new information may change this belief.

How sensitive is the timing of transition away from the variable-price contract to current and past prices?

Findings in brief:

- ▶ Households respond to history of prices by switching contracts
- ▶ Duration on contract vary substantially across alternative price histories
- ▶ Not necessarily the case that households interpret higher variable prices as evidence that a fixed-price contract is preferable

The decision problem

Household have to decide between two states of the world:

- ▶ Variable price \succ Fixed price
- ▶ Variable price \prec Fixed price

given all the information available at time t (time since start of contract):

- ▶ The history \mathcal{X}_t provides households with information about the state of the world (e.g., the sign and magnitude of the expected price differential and temperature)
- ▶ z_0 is household characteristics (e.g., initial characteristics)

Belief update

We assume that decision maker learns about the state of the world, S_t , in a Bayesian manner.

$$S_t = \begin{cases} 1 & \text{variable price preferable} \\ 0 & \text{fixed price preferable} \end{cases}$$

$$\underbrace{\frac{P[S_t = 1 | \mathcal{X}_t, t, z_0, o_0]}{P[S_t = 0 | \mathcal{X}_t, t, z_0, o_0]}}_{\text{posterior odds}} = \underbrace{\frac{P[S_t = 1 | \mathcal{X}_{t-1}, t-1, z_0, o_0]}{P[S_t = 0 | \mathcal{X}_{t-1}, t-1, z_0, o_0]}}_{\text{prior odds}} \cdot \underbrace{\frac{P[x_t | S_t = 1, \mathcal{X}_{t-1}, t]}{P[x_t | S_t = 0, \mathcal{X}_{t-1}, t]}}_{\text{likelihood ratio}}$$

$$\Rightarrow \ln o_t = \ln o_{t-1} + \ln \lambda_t$$

Modified Bayesian Updating

- ▶ Prior odds at time t are obtained as a simple transformation of the posterior odds obtained after observing all available information at time $t - 1$:

$$\underbrace{\frac{P[S_{t+1} = 1 | \mathcal{X}_t, t, z_0, o_0]}{P[S_{t+1} = 0 | \mathcal{X}_t, t, z_0, o_0]}}_{\text{prior odds at time } t+1} = \exp(\kappa_0) \left(\underbrace{\frac{P[S_t = 1 | \mathcal{X}_t, t, z_0, o_0]}{P[S_t = 0 | \mathcal{X}_t, t, z_0, o_0]}}_{\text{posterior odds at time } t} \right)^{\kappa_1}$$

$$\Rightarrow \ln o_t = \kappa_0 + \kappa_1 \ln o_{t-1} + \ln \lambda_t$$

- ▶ This expression describes how, starting from initial odds, the past information accumulates into today's odds depending on the parameters κ_0 and κ_1

Modeling the likelihood ratio λ_t

- ▶ λ_t measures how the information observed at time t contributes to the household's view about the state of the world.
- ▶ Simple AR(1) models with homoscedastic normal innovations. Decision maker assumes x_t , current information, evolves as:

$$x_t \sim \mathcal{N}(\rho_- x_{t-1} + \mu_-, \sigma_-^2) \text{ if } S = 1,$$

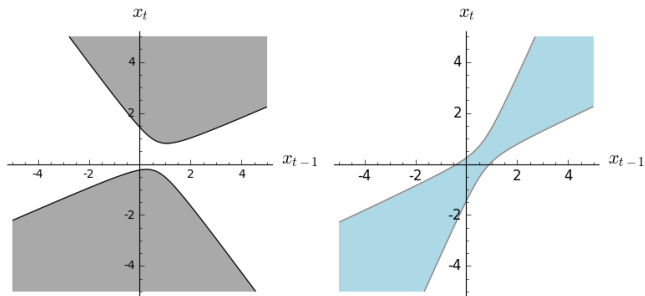
$$x_t \sim \mathcal{N}(\rho_+ x_{t-1} + \mu_+, \sigma_+^2) \text{ if } S = 0,$$

- ▶ Alternatively, keep AR(1) model, but allow state-specific variance:

$$x_t \sim \mathcal{N}(\rho_- x_{t-1} + \mu_-, \sigma_-^2) \text{ if } S = 1,$$

$$x_t \sim \mathcal{N}(\rho_+ x_{t-1} + \mu_+, \sigma_+^2) \text{ if } S = 0,$$

Figure 1: Sign of Log Likelihood Ratio



Note: The shaded area in each case indicates the region in the plane where the log likelihood ratio is positive. The parameters for the regions on the left hand pane are: $\mu_- = -0.1$, $\mu_+ = 0.1$, $\rho_- = 0.6$, $\rho_+ = 0.3$, $\sigma_- = 0.65$ and $\sigma_+ = 0.55$. To produce the left hand pane, the parameters are kept identical except $\sigma_- = 0.55$ and $\sigma_+ = 0.65$.

Probability to Keep/Change Contract

The final part of the model concerns the probability to keep the variable-price contract.

- ▶ We assume: $P[D_t = 1 | \mathcal{X}_t, t, z_0, o_0]$, $t = 1, \dots + \infty$, satisfies a logit specification which depends on the initial characteristics of the household, z_0 , and of the log-odds, $\ln o_t$:

$$P[D_t = 1 | \mathcal{X}_t, t, z_0, o_0] = \frac{1}{1 + \exp\{-z_0\pi_0 - \pi_1 \ln o_t\}}$$

where π_0 and π_1 are parameters.

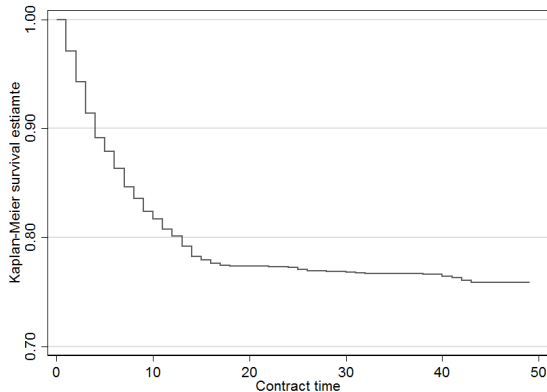
$\ln o_t$ depends on the current information summarized by $\ln \lambda_t$.

- ▶ The specification of the probabilities allow for a simple expressions for the individual contributions to the (conditional) likelihood.

Data

- ▶ Data on Skellefteå Krafts customers
- ▶ New customers (3353 households) starting variable-price contract June 2010 - February 2012. Sample period ends in June 2014
- ▶ 730 households switch to a fixed-price contract
- ▶ Contract choice, prices (seasonally adjusted)
- ▶ Housing type, heating system at household level
- ▶ Income, education, number of inhabitants at zip-code level
- ▶ Temperature at postal area level (SMHI, seasonally adjusted)

Figure 2: Kaplan-Meier survival estimate



Note: The figure shows the Kaplan-Meier estimator for the proportion of households remaining on variable-price contracts over time

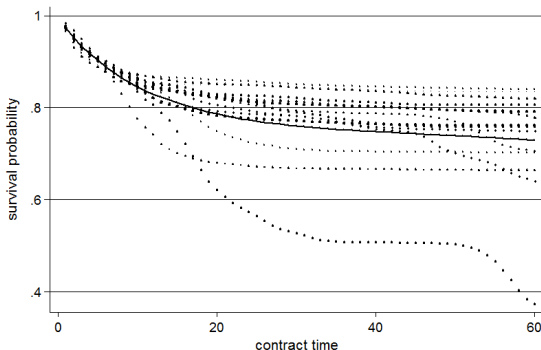
Estimation results

- ▶ Better model fit for heteroskedastic information with trend and discount (LL, AIC, BIC)
- ▶ Households downplay past information

Parameter estimates (price responsiveness) somewhat hard to interpret for heteroskedastic model (many interaction terms)

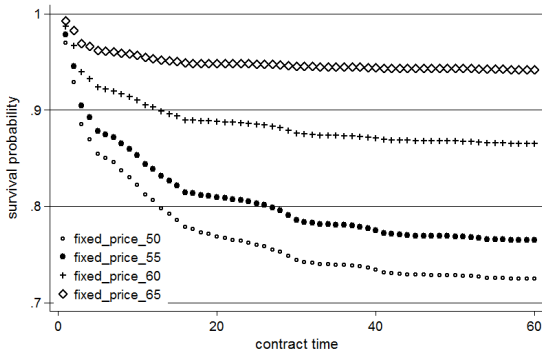
- ▶ A straight-forward way of understanding and illustrating results is to produce survival function(s):
- ▶ Estimate a time series model for the variable price (AR(1)-ARCH(1)), simulate alternative price differential histories;
- ▶ Given that history, calculate the survival probability given the parameter estimates from the survival model;

Average survival vs. history specific survival, no temperature, homoscedastic information



Note: The figure shows the "average" survival function (survival with a variable-price contract) against the survival function for specific variable price histories.

The average survival function of a variable-price tariff, temperature info, heteroscedastic information



Note: The figure shows the "average" survival function (survival with a variable-price contract) against different fixed prices. Average weather.

Conclusions

- ▶ Response to prices not obvious; depends on specification of the likelihood ratio
- ▶ Households respond to flow of information
- ▶ The response is seemingly small in the population, but may be bigger for alternative histories of prices
- ▶ Alternative price processes likely in the future, as share of intermittent generation increase
- ▶ If more households switch to fixed-price contracts, less price response in the short run